

RESEARCH LETTER

10.1002/2015GL067325

Key Points:

- The effect of cross-scale interactions in air temperature variability is quantified
- Seven to eight year cycle influences the amplitude of the annual cycle in Europe
- Seven to eight year cycle influences interannual variability of air temperature anomalies

Supporting Information:

- Text S1 and Figures S1–S8

Correspondence to:

M. Paluš,
mp@cs.cas.cz

Citation:

Jajcay, N., J. Hlinka, S. Kravtsov, A. A. Tsonis, and M. Paluš (2016), Time scales of the European surface air temperature variability: The role of the 7–8 year cycle, *Geophys. Res. Lett.*, 43, 902–909, doi:10.1002/2015GL067325.

Received 10 DEC 2015

Accepted 30 DEC 2015

Accepted article online 7 JAN 2016

Published online 23 JAN 2016

Time scales of the European surface air temperature variability: The role of the 7–8 year cycle

Nikola Jajcay^{1,2}, Jaroslav Hlinka¹, Sergey Kravtsov³, Anastasios A. Tsonis³, and Milan Paluš¹

¹Department of Nonlinear Dynamics and Complex Systems, Institute of Computer Science, Academy of Sciences of the Czech Republic, Prague, Czech Republic, ²Department of Atmospheric Physics, Faculty of Mathematics and Physics, Charles University, Prague, Czech Republic, ³Department of Mathematical Sciences, Atmospheric Science Group, University of Wisconsin-Milwaukee, Milwaukee, Wisconsin, USA

Abstract Air temperature variability on different time scales exhibits recurring patterns and quasi-oscillatory phenomena. Climate oscillations with the period about 7–8 years have been observed in many instrumental records in Europe. Although these oscillations are weak if considering their amplitude, they might have nonnegligible influence on temperature variability on shorter time scales due to cross-scale interactions recently observed by Paluš (2014). In order to quantify the cross-scale influence, we propose a simple conditional mean approach which estimates the effect of the cycle with the period close to 8 years on the amplitude of the annual cycle in surface air temperature (SAT) in the range 0.7–1.4°C and the effect on the overall variability of the SAT anomalies (SATA) leads to the changes 1.5–1.7°C in the annual SATA means. The strongest effect in the winter SATA means reaches 4–5°C in central European station and reanalysis data.

1. Introduction

The Earth climate, in general, and the air temperature, in particular, vary on many spatial and temporal scales. We will focus on a relatively small temporal range from the annual scale to near-decadal time scales, in long-term surface air temperature (SAT) records from the European midlatitudes. Long instrumental temperature records available from a number of European stations give a unique opportunity to study the long-term temperature variability and to identify repeating patterns such as interannual climate oscillations. Analyzing the 335 year long central England temperature (CET) record, *Plaut et al.* [1995] identified climate oscillations with the period about 7–8 years. *Baliunas et al.* [1997] and *Benner* [1999] have confirmed this finding in CET, while *Paluš and Novotná* [1998] detected an oscillatory phenomenon in the same frequency range in the 223 year long SAT record from Prague-Klementinum, as well as in SAT from other European locations, e.g., De Bilt, Berlin, and Wrocław [*Paluš and Novotná*, 2004]. *Pišoft et al.* [2004] and *Brázdil et al.* [2012] observed the 7–8 year cycle in SAT from various stations in the Czech Republic. *Grieser et al.* [2002] reported the 7–8 year oscillations in SAT records from western and northern Europe, and *Pišoft et al.* [2009] identified spatial patterns of occurrence of the 8 year cycle at various geopotential levels in the National Centers for Environmental Prediction/National Center for Atmospheric Research reanalyzed temperature series. *Sen and Ogrin* [2015] extended this list using SAT from Zagreb, while *Čermák et al.* [2014] detected the 7–8 year cycle in soil and ground temperatures collected in the Prague Geothermal Climate Change Observatory. The 7–8 year oscillations have also been observed in various climate-related data in Europe, the North Atlantic, and the Mediterranean regions [see, e.g., *Gámiz-Fortis et al.*, 2002; *Kondrashov et al.*, 2005; *Feliks et al.*, 2010]. These cycles have usually been observed using subtle detection techniques such as the singular spectrum analysis (SSA) [*Vautard et al.*, 1992], Monte Carlo SSA [*Allen and Smith*, 1996], or Enhanced Monte Carlo SSA [*Paluš and Novotná*, 1998, 2004] since their amplitude is typically very low and the cycles, e.g., in the air temperature, are hidden in the overall temperature variability. For example, in our estimates, described below, the amplitude of the cycle with the period about 8 years (the red curve in Figures 1a and 1b), extracted from the Prague-Klementinum SAT record, is smaller than 0.5°C. We can observe, however, approximately 8 year modulation of the amplitude of the annual cycle (the blue curve) and of the winter minima (Figure 1a, SAT data in grey) in markedly greater ranges. Therefore, it is important to understand possible relations between slow cycles and faster temperature variability.

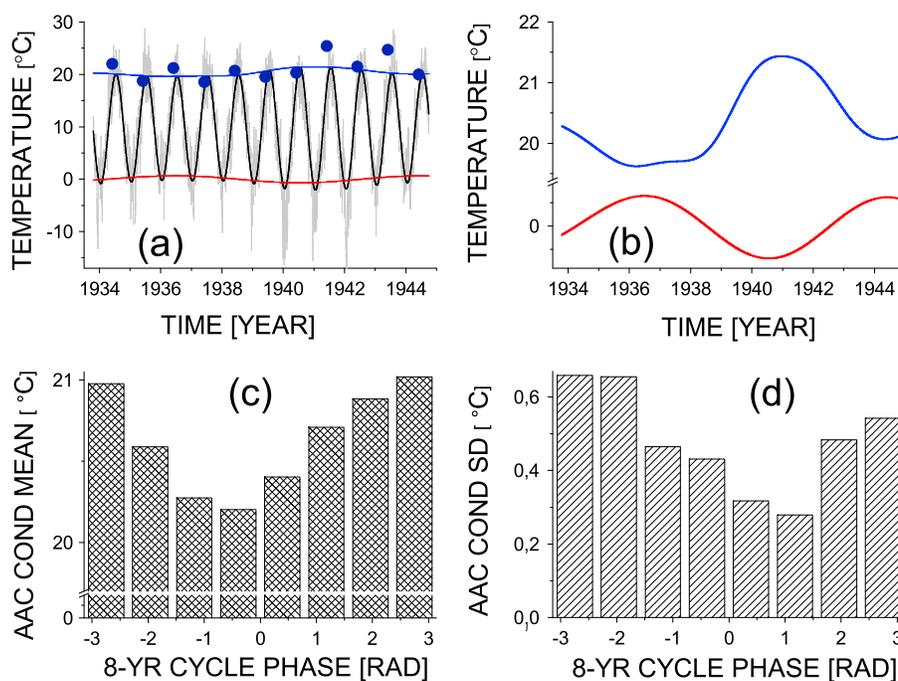


Figure 1. (a) Cycles in the SAT data in the period 20 October 1933 to 1 October 1944. Shown are the SAT daily station data from Prague-Klementinum (grey), the wavelet reconstruction $A_{8y}(t) \cdot \cos \phi_{8y}(t)$ of the 8 year cycle (red), the wavelet reconstruction of the annual cycle $A_{1y}(t) \cdot \cos \phi_{1y}(t)$ (black), the wavelet amplitude $A_{1y}(t)$ of the annual cycle (blue), and the climatological amplitude (dark blue dots, see the supporting information for details). (b) The wavelet reconstruction of the 8 year cycle (red) and the wavelet amplitude of the annual cycle (blue) “zoomed” in their individual scales. (c) Conditional means and (d) conditional standard deviations (SD) for the amplitude of the annual cycle, $A_{1y}(t)$, for the Prague-Klementinum SAT data within the period 1 January 1962 to 31 December 2009, conditioned on the phase $\phi_{8y}(t)$ of the 8 year cycle, divided into eight equidistant bins. Note that each bin represents approximately 1 year of the 8 year cycle.

Studying interactions between dynamics on various time scales in long-term daily SAT records from European locations, Paluš [2014] has observed an information transfer from larger to smaller time scales in the form of a causal influence (causality in the Granger sense [see Hlaváčková-Schindler et al., 2007]) of the phase of slow temperature oscillations on the amplitude of faster temperature variability. The influenced faster variability is characterized by the temporal scales from a few months to 4–5 years, while the periods of the influencing oscillatory phenomenon vary between 6 and 11 years; however, the most probable period is between 7 and 8 years [see Paluš, 2014, Figure 3a].

In this letter we propose an approach to quantify the effect of the cross-scale influence of the oscillatory mode with the period close to 8 years on the amplitude of the annual temperature cycle, as well as on the overall variability of SAT anomalies (SATA) in Europe. We will show that this effect is nonstationary, variable in space and time; however, it is nonnegligible and significantly larger than the amplitude of the 7–8 year cycle itself. Therefore, this phenomenon requires further research and understanding of its mechanisms and assessment of potential predictability skills related to interannual temperature variability.

2. Data and Methods

2.1. Data

Daily mean SAT data recorded at several stations with sufficiently long uninterrupted record provided by the ECA&D project [Klein Tank et al., 2002] were used for the study of temporal evolution of the effect of the cross-scale interactions in the temperature variability. Results from Prague-Klementinum (longitude 14°25'E, latitude 50°05'N), which record spans the period 1 January 1775 to 31 December 2013 are presented here. Results from two German stations, Hamburg-Fuhlsbüttel (9°59'E, 53°38'N) with the record spanning 1 January 1891 to 31 December 2013 and Potsdam (13°04'E, 52°23'N) with the record spanning 1 January 1893 to 31 December 2013, are presented in the supporting information.

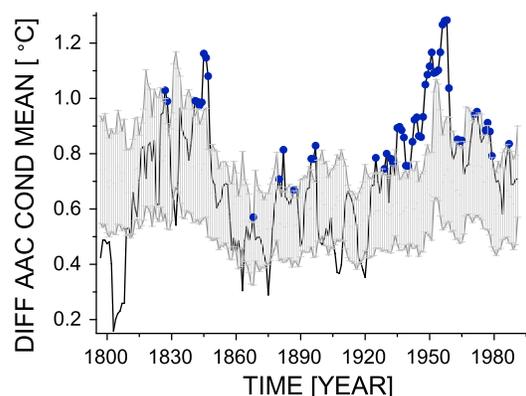


Figure 2. Temporal evolution of the effect of the 8 year cycle on the amplitude of the annual cycle in the Prague-Klementinum daily SAT. The differences between the maximum and minimum AAC conditional means (black curve), tested using 1000 FT surrogates (the means plotted as the grey curve; the 95th percentile of the surrogate distribution is plotted using the light grey curve, connected by the grey bars with the surrogate means). Windows with statistically significant differences are marked by the blue dots, plotted in the middle of the window of the effective length 36.86 years (see the supporting information for details).

using the continuous complex wavelet transform (CCWT) with the Morlet mother wavelet [Torrence and Compo, 1998] which provides the instantaneous amplitude $A_f(t)$ and the instantaneous phase $\phi_f(t)$ for each wavelet central frequency f . Details and other approaches are described, e.g., by Paluš et al. [2005].

The instantaneous phase of the oscillations with the period close to 8 years is estimated from daily (station or reanalysis) SATA data using the CCWT with the central wavelet frequency corresponding to the period 8 years. The same method using the central wavelet period of 1 year is also used to extract the instantaneous phase and amplitude of the annual cycle from the SAT data. As the continuous wavelet transform results in a redundant decomposition, the real part $A_f(t) \cdot \cos \phi_f(t)$ of the wavelet component representing the annual cycle is regressed to fit the SAT data in order to interpret the amplitude values in °C.

Conditional means of either the amplitude of the annual cycle or the overall SATA variability, conditioned on the phase of the 8 year cycle, are computed using a simple binning technique. The phase interval $(-\pi, \pi)$ (representing the full cycle) is divided into eight bins. For each bin we evaluate the mean (or standard deviation) and obtain a discretized estimate of the conditional mean of the studied variable. See the supporting information for details. If the 8 year cycle has no influence on the studied variable, the conditional means in all eight bins should be the same, equal to the unconditional, global mean. Or, due to the finite sample effect, the conditional means should randomly fluctuate around the unconditional mean. If the conditional means vary as a function of the phase of the 8 year cycle, we will use the difference between the maximum and minimum conditional mean as a measure of the effect of the 8 year cycle on the studied variable. The statistical significance of such an effect is evaluated using the surrogate data method. The Fourier transform (FT) surrogate data [Theiler et al., 1992] represent a null hypothesis of a process with the same power spectrum as the studied data, but without any cross-frequency interaction. The rejection of this null hypothesis provides a statistical evidence for a cross-scale effect of the 8 year cycle on temperature variability on shorter time scales. See the supporting information for details and other surrogate data types.

3. Amplitude of the Annual Cycle in SAT

The strongest mode of variability in the European temperature data is the annual cycle. Its amplitude, however, varies in time and space [see, e.g., Zveryaev, 2007]. The wavelet reconstruction $\psi_{1y}(t) = A_{1y}(t) \cdot \cos \phi_{1y}(t)$ (the black curve in Figure 1a), regressed to match the SAT data (grey) is plotted in Figure 1a along with the amplitude $A_{1y}(t)$ of the annual cycle (blue) and the reconstruction $\psi_{8y}(t) = A_{8y}(t) \cdot \cos \phi_{8y}(t)$ (red) of the 8 year cycle.

For the study of the spatial variability of the effect over Europe, the temperature reanalysis data from E-OBS data set from the EU-FP6 project ENSEMBLES provided by the ECA&D project [Haylock et al., 2008] were used. The data set has daily temporal and $0.25^\circ \times 0.25^\circ$ spatial resolution. The spatial domain is confined to the range from 35°N to 65°N and from 12.5°W to 40°E . The temporal domain spans the period of 1 January 1950 to 31 December 2013.

2.2. Methods

We use the raw SAT data when the variability of the annual cycle is studied. The SAT anomalies (SATA), obtained from the SAT data by subtracting the average annual cycle (see the supporting information for details), are used for analyses of the variability on all other temporal scales (overall variability thereafter).

An oscillatory signal $s(t)$ can be conveniently described using its instantaneous amplitude $A(t)$ and phase $\phi(t)$. A broadband signal recorded from a multiscale process needs to be filtered into a spectral band of interest. In this study, both the filtering and amplitude and phase estimation is performed

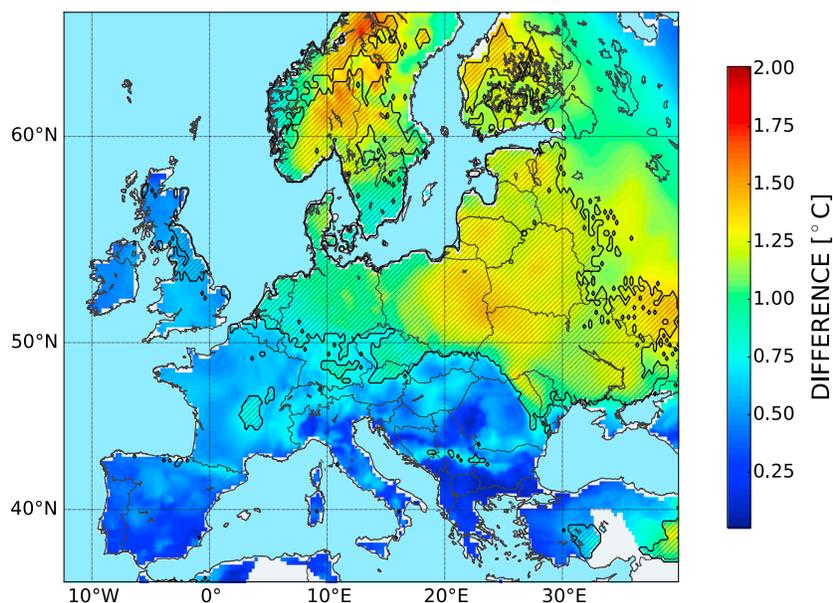


Figure 3. Spatial variability of the effect of the 8 year cycle on the amplitude of the SAT annual cycle in Europe. Differences of the maximum and minimum conditional means of the ECA&D reanalysis SAT annual cycle amplitude, $A_{1y}(\mathbf{x}, t)$, conditioned on the phase of the 8 year cycle, $\phi_{8y}(\mathbf{x}, t)$. The hatch pattern marks the areas where the effect is statistically significant.

A “climatological amplitude” (dark blue dots in Figure 1a), defined as the difference between the means of daily temperatures above the upper quartile and below the lower quartile in each year, is in a good agreement with the amplitude $A_{1y}(t)$ acquired from the wavelet transform.

The apparent relationship (visually enhanced using individual scales in Figure 1b) between the amplitude of the annual cycle (AAC thereafter) and the 8 year cycle (Pearson correlation coefficient -0.86) is further studied using the conditional mean technique, conditioning the AAC $A_{1y}(t)$ means on the phase $\phi_{8y}(t)$ of the 8 year cycle. The histogram of the conditional means of the AAC is presented in Figure 1c. The maximum mean, conditioned on the phase of the 8 year cycle, is located in the eighth bin at 21.02°C , while the minimum is located in the fourth bin at 20.20°C . This implies that through the 8 year cycle, the AAC changes, on average, within the range of 0.82°C . The conditional standard deviations of the AAC are ranging from 0.28 to 0.66°C with higher values in the outer bins and lower values in the central bins (Figure 1d).

The change 0.82°C of the AAC within the 8 year cycle is the average change for the six cycles in the period 1 January 1962 to 31 December 2009. When using different segments of the data, the results differ due to nonstationarity of the temperature data and their cross-scale interactions. In order to depict this nonstationarity, the temporal evolution of the difference between the maximum and minimum conditional means was characterized using the sliding window of 16,384 daily SAT samples (Figure 2). In the 239 year long Prague-Klementinum record the differences range from 0.2 to 1.3°C and they are statistically significant mainly during the last 90 years. The long-term evolution of the AAC variability reflects a modulation on the scale 60–80 years (Figure 2). This is the scale of the Atlantic Multidecadal Oscillation which influences large-scale climate variability modes [Wyatt *et al.*, 2012] and SAT variability on global [Chylek *et al.*, 2014a] and regional [Chylek *et al.*, 2014b] scales. However, inclusion of the multidecadal scales into the study of the cross-scale interactions is left for future research.

The gridded temperature reanalysis data on the dense $0.25^{\circ} \times 0.25^{\circ}$ spatial grid underwent the same AAC conditional means analysis using the phase of the CCWT-extracted 8 year cycle. The areas with the statistically significant maximum differences of the AAC conditional means are marked by the hatch pattern in Figure 3. The marked influence of the phase of the 8 year oscillatory mode on the amplitude of the annual cycle can be seen over central, northern, and eastern Europe.

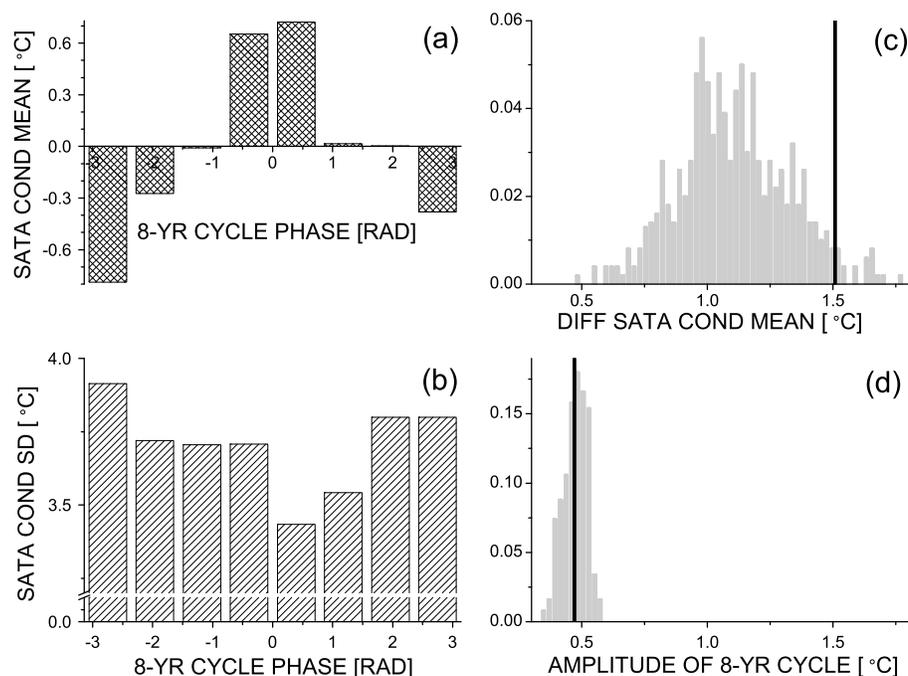


Figure 4. (a) Conditional means and (b) conditional standard deviations for the daily SATA from Prague-Klementinum in the period 1 January 1962 to 31 December 2009 conditioned on the phase of the 8 year cycle, $\phi_{8y}(t)$. (c) The difference between the maximum and minimum SATA conditional means depicted by the position of the black bar ($\approx 1.5^\circ\text{C}$) and the distribution of the same differences obtained from 1000 realizations of the FT surrogate data (grey histogram). (d) The mean amplitude of the 8 year cycle depicted by the position of the black bar ($<0.5^\circ\text{C}$) and the distribution of the mean amplitudes of the 8 year cycles obtained from 1000 realizations of the FT surrogate data (grey histogram).

4. Overall SATA Variability in the 8 Year Cycle

Above we have examined the effect of the 8 year cycle on the amplitude of the annual cycle. Since the 8 year cycle has an effect on various temporal scales [Paluš, 2014], now we explore its effect on the overall variability represented by the surface air temperature anomalies. Again, the conditional mean technique with the eight phase bins is used and the results for the SATA data from Prague-Klementinum are presented in Figure 4. The conditional SATA means, conditioned on the phase of the 8 year cycle (Figure 4a), show that negative temperature anomalies (“cold bins”) prevail at the beginning and the end of the 8 year cycle (with the minimum -0.79°C in the first bin), while the positive anomalies (“warm bins”) dominate the middle of the cycle (with the warmest bin having 0.72°C anomaly). The conditional standard deviations (Figure 4b) are behaving in the opposite way.

The difference between the maximum and minimum SATA conditional means, i.e., the effect of the 8 year cycle on the SATA variability, is approximately 1.5°C . This value exceeds the 98th percentile of the related surrogate data distribution (Figure 4c); i.e., it is statistically significant with $p < 0.02$. The amplitude of the 8 year cycle, estimated from the SATA data, is less than 0.5°C and is well reproduced in the FT surrogate data (Figure 4d). These computational statistics (Figures 4c and 4d) support the hypothesis that the 1.5°C difference in the SATA conditional means is not a result of random variability, neither can be explained by an 8 year component, linearly added to a background variability. Since the amplitude of the 8 year cycle itself is smaller than 0.5°C , the difference 1.5°C of the annual means during the 8 year cycle is mainly the result of the cross-scale interactions of the 8 year cycle with the variability on shorter time scales.

The effect of the 8 year cycle on the overall SATA variability, similarly as its effect on the amplitude of the annual cycle, varies in time and space. It also differs when we consider different seasons. The strongest effect was observed for the winter (December–February, DJF) season, when the differences between the maximum and minimum mean winter temperature, conditioned on the phase of the 8 year cycle, reach approximately $4\text{--}5^\circ\text{C}$ in the station SATA data from central Europe. During the summer season the effect is not statistically significant; i.e., it is not distinguishable from random variability. See the supporting information for details.

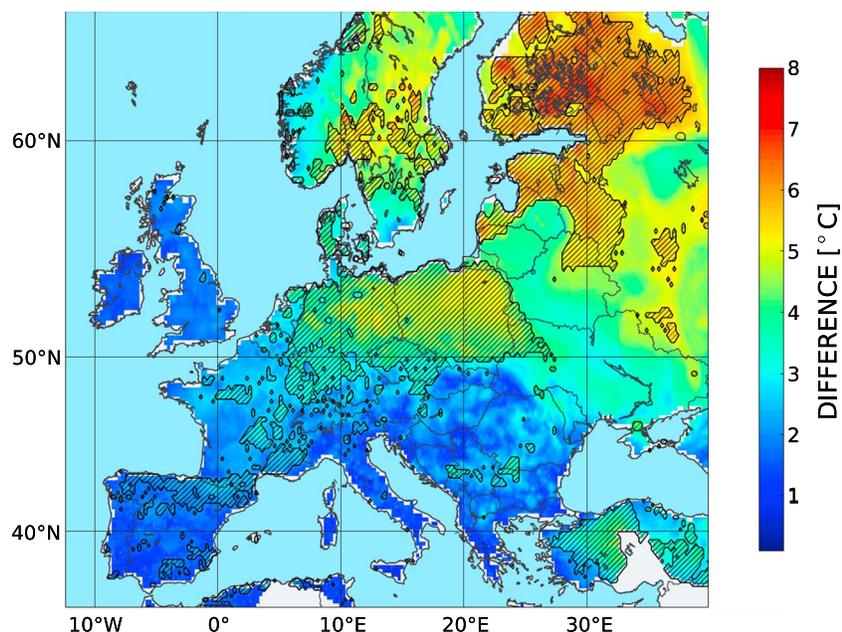


Figure 5. Spatial variability of the effect of the 8 year cycle on the mean winter temperature in Europe. Differences of the maximum and minimum conditional DJF means of the ECA&D reanalysis SATA, conditioned on the phase of the 8 year cycle, $\phi_{8y}(\mathbf{x}, t)$. The hatch pattern marks the areas where the effect is statistically significant.

The spatial variability of the effect on the winter temperature in Europe, estimated from the ECA&D SATA data, is presented in Figure 5. The areas with the statistically significant maximum differences of the conditional SATA means in the DJF season are marked by the hatch pattern. The differences range from about 1°C in Spain to the maximum of 6.5°C in Finland and adjacent areas of Russia. The pattern is similar to that in Figure 3, and in central and eastern Europe they both resemble (the inverse of) the mountain topography: the effect of the 8 year cycle is strong in the lowlands from the North and Baltic Seas southward and weakens at the mountain ranges of the Alps and Carpathians. The interaction of variable jet stream with the mountain topography apparently has a modulating effect on the influence of the 8 year cycle on the temperature variability in Europe.

5. Discussion and Conclusions

Considering air temperature variability in a range of time scales, *Paluš* [2014] presented a statistical evidence for a cross-scale-directed information flow from larger to smaller time scales in long-term SAT records from European stations. The phase of a slow oscillatory process influences temperature variability on shorter time scales. The influencing oscillatory phenomenon has variable frequency; however, its most probable period is close to 8 years, and for this period it has also the strongest effect [see *Paluš*, 2014, Figures 2b and 3a]. These periods (time scales) are consistent with the observations of an oscillatory mode with the period between 7 and 8 years in long-term temperature and other meteorological records in Europe (see the references in section 1). Therefore, in this study we have applied a simple conditional mean technique in order to quantitatively estimate the effect of the oscillatory mode with the period close to 8 years on the surface air temperature variability in Europe. The cycle itself has a small amplitude (<0.5°C in the presented example of the station SAT record from Prague-Klementinum, see Figure 4d) and is hidden in overall temperature variability. However, due to the cross-scale interactions [*Paluš*, 2014], the 8 year cycle influences the temperature variability on shorter time scales. The amplitude of the annual cycle in SAT changes within this cycle by 0.7–1.4°C, and the overall variability of SAT anomalies changes in annual means by 1.5–1.7°C. The strongest effect of the 8 year cycle has been observed in the winter season—the DJF SATA means change in the range 4–5°C. (This summary is restricted to twentieth century central Europe, where we have available both the station and reanalysis data giving consistent results.) These results suggest that the weak 7–8 year cycle plays a very important role in the temperature variability on interannual and shorter time scales. Therefore, this phenomenon deserves further study and understanding of its mechanisms.

Paluš [2014] hypothesizes that in the analyzed SAT data, we have observed a regional manifestation of a general phenomenon of cross-scale interactions in the atmospheric dynamics in which global, low-frequency modes influence local, high-frequency variability. For instance, *Chekroun et al.* [2011] reported that the phase of the low-frequency modes of the El Niño–Southern Oscillation influences high-frequency variability (“weather noise”) of the sea surface temperature in the tropical Pacific. For the data analyzed in this study, the most relevant global mode is probably the North Atlantic Oscillation (NAO). The influence of the NAO on the air temperature in Europe is known [*Marshall et al.*, 2001], and its mechanisms depending on the phase of the NAO are described, e.g., by *Hurrell and Dickson* [2005]. Typically, Pearson’s correlations have been computed between (mostly winter) air temperature records and NAO indices [see, e.g., *Pokorná and Huth*, 2015]; however, specific time scales have not been considered yet, although the 7–8 year cycle has also been detected in the NAO index [*Gámiz-Fortis et al.*, 2002; *Paluš and Novotná*, 2004]. Our results demonstrate the importance of understanding the climate variability in scale-specific regional modes and their cross-scale interactions and causal relations with global circulation variability modes which are localized not only in space [*Vejmelka et al.*, 2014] but also in a time scale or in a frequency range [*Groth and Ghil*, 2011, 2015].

Acknowledgments

The data used are listed in the references. This study was supported by the Ministry of Education, Youth and Sports of the Czech Republic within the Program KONTAKT II, project LH14001 (N.J. and M.P.) and also by NSF grants OCE-1243158 (S.K.) and AGS-1408897 (S.K. and A.A.T.).

References

- Allen, M. R., and L. A. Smith (1996), Monte Carlo SSA: Detecting irregular oscillations in the presence of colored noise, *J. Clim.*, *9*(12), 3373–3404, doi:10.1175/1520-0442(1996)009<3373:MCSPIO>2.0.CO;2.
- Baliunas, S., P. Frick, D. Sokoloff, and W. Soon (1997), Time scales and trends in the central England temperature data (1659–1990): A wavelet analysis, *Geophys. Res. Lett.*, *24*(11), 1351–1354, doi:10.1029/97GL01184.
- Benner, T. C. (1999), Central England temperatures: Long-term variability and teleconnections, *Int. J. Climatol.*, *19*(4), 391–403, doi:10.1002/(SICI)1097-0088(19990330)19:4<391::AID-JOC365>3.0.CO;2-Z.
- Brázdil, R., P. Zahradníček, P. Pišoft, P. Štěpánek, M. Bělínová, and P. Dobrovolný (2012), Temperature and precipitation fluctuations in the Czech Republic during the period of instrumental measurements, *Theor. Appl. Climatol.*, *110*(1–2), 17–34, doi:10.1007/s00704-012-0604-3.
- Čermák, V., L. Bodri, J. Šafanda, M. Krešl, and P. Dědeček (2014), Ground-air temperature tracking and multi-year cycles in the subsurface temperature time series at geothermal climate-change observatory, *Stud. Geophys. Geod.*, *58*(3), 403–424, doi:10.1007/s11200-013-0356-2.
- Chekroun, M. D., D. Kondrashov, and M. Ghil (2011), Predicting stochastic systems by noise sampling, and application to the El Niño–Southern Oscillation, *Proc. Natl. Acad. Sci. U.S.A.*, *108*(29), 11,766–11,771, doi:10.1073/pnas.1015753108.
- Chylek, P., J. D. Klett, G. Lesins, M. K. Dubey, and N. Hengartner (2014a), The Atlantic Multidecadal Oscillation as a dominant factor of oceanic influence on climate, *Geophys. Res. Lett.*, *41*, 1689–1697, doi:10.1002/2014GL059274.
- Chylek, P., M. K. Dubey, G. Lesins, J. Li, and N. Hengartner (2014b), Imprint of the Atlantic Multi-Decadal Oscillation and Pacific Decadal Oscillation on southwestern US climate: Past, present, and future, *Clim. Dyn.*, *43*(1–2), 119–129, doi:10.1007/s00382-013-1933-3.
- Feliks, Y., M. Ghil, and A. W. Robertson (2010), Oscillatory climate modes in the Eastern Mediterranean and their synchronization with the North Atlantic Oscillation, *J. Clim.*, *23*(15), 4060–4079, doi:10.1175/2010JCLI3181.1.
- Gámiz-Fortis, S., D. Pozo-Vázquez, M. Esteban-Parra, and Y. Castro-Diez (2002), Spectral characteristics and predictability of the NAO assessed through Singular Spectral Analysis, *J. Geophys. Res.*, *107*(D23), 4685, doi:10.1029/2001JD001436.
- Grieser, J., S. Trömel, and C.-D. Schönwiese (2002), Statistical time series decomposition into significant components and application to European temperature, *Theor. Appl. Climatol.*, *71*(3–4), 171–183, doi:10.1007/s007040200003.
- Groth, A., and M. Ghil (2011), Multivariate singular spectrum analysis and the road to phase synchronization, *Phys. Rev. E*, *84*, 036206, doi:10.1103/PhysRevE.84.036206.
- Groth, A., and M. Ghil (2015), Monte Carlo Singular Spectrum Analysis (SSA) revisited: Detecting oscillator clusters in multivariate datasets, *J. Clim.*, *28*, 7873–7893, doi:10.1175/JCLI-D-15-0100.1.
- Haylock, M. R., N. Hofstra, A. M. G. Klein Tank, E. J. Klok, P. D. Jones, and M. New (2008), A European daily high-resolution gridded data set of surface temperature and precipitation for 1950–2006, *J. Geophys. Res.*, *113*, D20119, doi:10.1029/2008JD010201.
- Hlaváčková-Schindler, K., M. Paluš, M. Vejmelka, and J. Bhattacharya (2007), Causality detection based on information-theoretic approaches in time series analysis, *Phys. Rep.*, *441*(1), 1–46, doi:10.1016/j.physrep.2006.12.004.
- Hurrell, J., and R. Dickson (2005), Climate variability over the North Atlantic, in *Marine Ecosystems and Climate Variation: The North Atlantic—A Comparative Perspective*, edited by N. C. Stenseth et al., pp. 15–314, Oxford Univ. Press, Oxford.
- Klein Tank, A. M. G., et al. (2002), Daily dataset of 20th-century surface air temperature and precipitation series for the European Climate Assessment, *Int. J. Climatol.*, *22*(12), 1441–1453, doi:10.1002/joc.773.
- Kondrashov, D., Y. Feliks, and M. Ghil (2005), Oscillatory modes of extended Nile River records (A.D. 622–1922), *Geophys. Res. Lett.*, *32*, L10702, doi:10.1029/2004GL022156.
- Marshall, J., Y. Kushnir, D. Battisti, P. Chang, A. Czaja, R. Dickson, J. Hurrell, M. McCartney, R. Saravanan, and M. Visbeck (2001), North Atlantic climate variability: Phenomena, impacts and mechanisms, *Int. J. Climatol.*, *21*(15), 1863–1898, doi:10.1002/joc.693.
- Paluš, M. (2014), Multiscale atmospheric dynamics: Cross-frequency phase-amplitude coupling in the air temperature, *Phys. Rev. Lett.*, *112*, 078702, doi:10.1103/PhysRevLett.112.078702.
- Paluš, M., and D. Novotná (1998), Detecting modes with nontrivial dynamics embedded in colored noise: Enhanced Monte Carlo SSA and the case of climate oscillations, *Phys. Lett. A*, *248*(2), 191–202.
- Paluš, M., and D. Novotná (2004), Enhanced Monte Carlo singular system analysis and detection of period 7.8 years oscillatory modes in the monthly NAO index and temperature records, *Nonlinear Processes Geophys.*, *11*(5–6), 721–729, doi:10.5194/npg-11-721-2004.
- Paluš, M., D. Novotná, and P. Tichavský (2005), Shifts of seasons at the European mid-latitudes: Natural fluctuations correlated with the North Atlantic Oscillation, *Geophys. Res. Lett.*, *32*, L12805, doi:10.1029/2005GL022838.
- Pišoft, P., J. Kalvová, and R. Brázdil (2004), Cycles and trends in the Czech temperature series using wavelet transforms, *Int. J. Climatol.*, *24*(13), 1661–1670, doi:10.1002/joc.1095.
- Pišoft, P., J. Mikšovský, and M. Žák (2009), An analysis of the spatial distribution of approximate 8 years periodicity in NCEP/NCAR and ERA-40 temperature fields, *Eur. Phys. J. Spec. Top.*, *174*(1), 147–155, doi:10.1140/epjst/e2009-01097-3.

- Plaut, G., M. Ghil, and R. Vautard (1995), Interannual and interdecadal variability in 335 years of central England temperatures, *Science*, 268(5211), 710–713, doi:10.1126/science.268.5211.710.
- Pokorná, L., and R. Huth (2015), Climate impacts of the NAO are sensitive to how the NAO is defined, *Theor. Appl. Climatol.*, 119(3–4), 639–652, doi:10.1007/s00704-014-1116-0.
- Sen, A. K., and D. Ogrin (2015), Analysis of monthly, winter, and annual temperatures in Zagreb, Croatia, from 1864 to 2010: The 7.7-year cycle and the North Atlantic Oscillation, *Theor. Appl. Climatol.*, doi:10.1007/s00704-015-1388-z.
- Theiler, J., S. Eubank, A. Longtin, B. Galdrikian, and J. Doyne Farmer (1992), Testing for nonlinearity in time series: The method of surrogate data, *Physica D*, 58(1), 77–94, doi:10.1016/0167-2789(92)90102-S.
- Torrence, C., and G. P. Compo (1998), A practical guide to wavelet analysis, *Bull. Am. Meteorol. Soc.*, 79(1), 61–78, doi:10.1175/1520-0477(1998)079<0061:APGTWA>2.0.CO;2.
- Vautard, R., P. Yiou, and M. Ghil (1992), Singular-spectrum analysis: A toolkit for short, noisy chaotic signals, *Physica D*, 58(1), 95–126, doi:10.1016/0167-2789(92)90103-T.
- Vejmelka, M., L. Pokorná, J. Hlinka, D. Hartman, N. Jajcay, and M. Paluš (2014), Non-random correlation structures and dimensionality reduction in multivariate climate data, *Clim. Dyn.*, 44(9–10), 2663–2682, doi:10.1007/s00382-014-2244-z.
- Wyatt, M. G., S. Kravtsov, and A. A. Tsonis (2012), Atlantic Multidecadal Oscillation and Northern Hemisphere's climate variability, *Clim. Dyn.*, 38(5–6), 929–949, doi:10.1007/s00382-011-1071-8.
- Zveryaev, I. (2007), Climatology and long-term variability of the annual cycle of air temperature over Europe, *Russ. Meteorol. Hydrol.*, 32(7), 426–430, doi:10.3103/S1068373907070023.